

1-1-1998

A Berry-Esseen Bound for Empty Boxes Statistic on the Scheme an Allocations of Several Type Balls

S.A. Mirakhmedov

O. Saidova

Follow this and additional works at: <https://journals.tubitak.gov.tr/math>



Part of the [Mathematics Commons](#)

Recommended Citation

Mirakhmedov, S.A. and Saidova, O. (1998) "A Berry-Esseen Bound for Empty Boxes Statistic on the Scheme an Allocations of Several Type Balls," *Turkish Journal of Mathematics*: Vol. 22: No. 1, Article 4. Available at: <https://journals.tubitak.gov.tr/math/vol22/iss1/4>

This Article is brought to you for free and open access by TÜBİTAK Academic Journals. It has been accepted for inclusion in Turkish Journal of Mathematics by an authorized editor of TÜBİTAK Academic Journals. For more information, please contact academic.publications@tubitak.gov.tr.

A BERRY-ESSEEN BOUND FOR EMPTY BOXES STATISTIC ON THE SCHEME AN ALLOCATIONS OF SEVERAL TYPE BALLS*

S.A. Mirakhmedov & O. Saidova

Abstract

A Berry-Esseen bound for the number of empty cells in the scheme of independent and random allocation of balls of s type into different cells is obtained.

Key words and phrases: Central limit theorem, empty cells, random allocations.

Introduction

Let n_1 balls be of first type and n_2 be balls of a second type, etc., and n_s balls of s th.type be distributed independently and randomly into N different cells, in such a way that each ball of i th type has probability p_{ik} of landing into k th cell, $p_{i1} + \dots + p_{iN} = 1$, $i = 1, \dots, s$. Let $\mu_0(s) = \mu_0(s, N, n_1, \dots, n_s)$ be a number of empty cells after all n_1, \dots, n_s losses. If $s = 1$ we deal with multinomial scheme of an allocation and $\mu_0(1)$ is a well-known empty box test statistic (see, for example, Koichin, Sevastjanov, Chistjacov (1976)). For example, the random variable (r.v.) $\mu_0(s)$ used as test statistic for verification of homogeneity hypothesis.

Here we get a bound for remainder in the central limit theorem for $\mu_0(s)$. Our theorem generalizes the result of Quine and Robinson (1982).

Result. We consider the case that s is fixed and $N = N(n_1, \dots, n_s)$ is growing as one of n_1, \dots, n_s increases. Suppose that for all $j = 1, \dots, s$ and $k = 1, \dots, N$

$$Np_{jk} \leq C_0 \quad \text{and} \quad n_i \leq \exp\{C_1 N\}. \quad (1)$$

Here and in what follows, $C_j, C_j(\cdot)$ are positive constants not dependent on N, n_1, \dots, n_s .

Denote

$$\lambda_{jm} = n_j p_{jm}, \quad \lambda_m^{(\epsilon)} = \lambda_{1m} + \dots + \lambda_{sm}, \quad \alpha_i = n_i/N,$$

* Research supported by TUBİTAK, TURKEY. We are grateful to Hacettepe University, whose hospitality we enjoyed while working on part of this paper.

$$\begin{aligned}
 A_N(s) &= \sum_{m=1}^N \exp\{-\lambda_m\}, \quad a_e = \frac{1}{n_e} \sum_{m=1}^N \lambda_{em} \exp\{-\lambda_m\}, \\
 \sigma_N^2 &= \sum_{m=1}^N \left[\exp\{-\lambda_m\}(1 - \exp\{-\lambda_m\}) - \sum_{j=1}^s \alpha_j a_j^2 \right], \\
 \omega_N^{(s)}(x) &= \left| P \left\{ \frac{\mu_0(s) - A_N(s)}{\sigma_N(s)} < x \right\} - \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x \exp\left\{-\frac{u^2}{2}\right\} \right|.
 \end{aligned}$$

Theorem. Under condition (1) for arbitrary $\nu \geq 0$ there exist $C(s, C_0, \nu)$ such that

$$\omega_N^{(s)}(x) \leq \frac{C(s, C_0, \nu)}{1 + |x|^\nu} \left[\frac{1}{\sigma_N(s)} + \sum_{j=1}^s \frac{1}{\sqrt{n_j}} \right].$$

Corollary 1. Under condition (1) there exist $C(s) > 0$ such that

$$\sup_x \omega_N^{(s)}(x) \leq C(s) \left[\frac{1}{\sigma_N(s)} + \sum_{j=1}^s \frac{1}{\sqrt{n_j}} \right].$$

Corollary 2. Suppose that $N, n_1, \dots, n_s \rightarrow \infty$ in such a way that $\sigma_N(s) \rightarrow \infty$ and (1) is hold true. Then $\mu_0(s)$ is asymptotically normal.

The result of Quine and Robinson (1982) is $\omega_N^{(1)}(x) \leq C\sigma_N^{-1}(1)$, which also follows from our theorem since $\sigma_N^2(1) \leq n_1$: but in the general case we have $\sigma_N^2(s) \leq n_1 + \dots + n_s$.

Proof. It is clear that

$$\mu_0(s) = \sum_{m=1}^N f(\eta_{1m}, \dots, \eta_{sm}),$$

where η_{ik} is a number of balls of i th type in m th cell after n_1, \dots, n_s tosses, and $f(0, \dots, 0) = 1$ and $f(y_1, \dots, y_s) = 0$ if $y_i > 0$ for some $i = 1, \dots, s$.

Let ζ_{jm} be a Poisson with parameter λ_{jm} , $\zeta_m^{(s)} = (\zeta_{1m}, \dots, \zeta_{sm})$, then

$$g_m(\zeta_m^{(s)}) = f(\zeta_m^{(s)}) - \exp\{-\lambda_m\} + \sum_{i=1}^s a_i(\zeta_{im} - \lambda_{im}).$$

From Corollary 2 of Mirakhmedov (1987) we get

$$\omega_N^{(s)}(x) \leq \frac{C(s, k)}{1 + |x|^k} \left[\beta_{3N} + \beta_{k+2, N} + \sum_{i=1}^s \frac{1}{\sqrt{n_i}} \right] \quad (2)$$

for any integer $k > 0$, where

$$\beta_{kN} = \frac{1}{\sigma_N^k(s)} \sum_{m=1}^N E|g_m(\zeta_m^{(s)})|^k.$$

We remark that

$$\sigma_N^2(s) = \sum_{m=1}^N Dg_m(\zeta_m^{(s)}).$$

We rewrite $\sigma_N^2(s)$ and $g_m(\zeta_m^{(s)})$ as follows:

$$\sigma_N^2(s) = \sum_{m=1}^N [1 - (1 + \lambda_m) \exp\{-\lambda_m\}] \exp\{-\lambda_m\} + \sum_{m=1}^N \sum_{j=1}^s \lambda_{jm} (\exp\{-\lambda_m\} - a_j)^2, \quad (3)$$

$$g_m(\zeta_m^{(s)}) = f(\zeta_m^{(s)}) + \exp\{-\lambda_m\} \sum_{i=1}^s \zeta_{im} - (1 + \lambda_m) \exp\{-\lambda_m\} + \sum_{i=1}^s (a_i - \exp\{-\lambda_m\})(\zeta_{im} - \lambda_{im}).$$

Then for arbitrary $b > 1$ we get

$$\begin{aligned} E|g_m(\zeta_m^{(s)})|^b &\leq 2^{b-1} E \left| f(\zeta_m^{(s)}) + \exp\{-\lambda_m\} \sum_{i=1}^s \zeta_{im} - (1 + \lambda_m) \exp\{-\lambda_m\} \right|^b \\ &\quad + (2s)^{b-1} \sum_{i=1}^s |a_i - \exp\{-\lambda_m\}|^b E|\zeta_{im} - \lambda_m|^b \\ &\equiv 2^{b-1} \Delta_{1m} + (2s)^{b-1} \Delta_{2m}. \end{aligned} \quad (4)$$

The r.v. $f(\zeta_m^{(s)})$ has the same distribution as r.v. $\varphi(\zeta_{1m} + \dots + \zeta_{sm})$ where $\varphi(0) = 1$ and $\varphi(x) = 0$ if $x > 0$. Thus

$$\begin{aligned} \Delta_{1m} &= E \left| \varphi(\zeta_{1m} + \dots + \zeta_{sm}) + \exp\{-\lambda_m\} \sum_{i=1}^s -(1 + \lambda_m) \exp\{-\lambda_m\} \right|^b \\ &\leq \exp\{-\lambda_m\} (1 - \exp\{-\lambda_m\} (1 + \lambda_m))^b + \lambda_m^{b+1} \exp\{-(b-1)\lambda_m\} \\ &\quad + \sum_{j=2}^{\infty} |j-1 - \lambda_m|^b \exp\{-(b+1)\lambda_m\} \frac{\lambda_m^j}{j!} \equiv \Delta'_{1m} + \Delta''_{1m} + \Delta'''_{1m} \end{aligned} \quad (5)$$

because $\zeta_{1m} + \dots + \zeta_{sm}$ is Poisson with parameter λ_m . Since $(1+u)e^{-u} < 1$ for $u > 0$ and (3), we have

$$\sum_{m=1}^N \Delta'_{1m} \leq \sum_{m=1}^N \exp\{-\lambda_m\} (1 - \exp\{-\lambda_m\} (1 + \lambda_m)) \leq \sigma_N^2(s). \quad (6)$$

From $u^2e^{-2u} \leq 1$ and $\frac{1}{2}u^2e^{-u} \leq 1 - (1+u)e^{-u}$, we get

$$\sum_{m=1}^N \Delta''_{1m} \leq \sum_{m=1}^N \lambda_m^2 \exp\{-2\lambda_m\} \leq 2 \sum_{m=1}^N (1 - (1 + \lambda_m)) \exp\{-\lambda_m\} \leq 2\sigma_N^2(s). \quad (7)$$

Let b be odd, $\sum_{\lambda_m \leq 1}$ and $\sum_{\lambda_m \geq 1}$ be a sum on m such that $\lambda_m \leq 1$ and $\lambda_m \geq 1$, correspondingly. We have

$$\begin{aligned} \sum_{\lambda_m \leq 1} \Delta'''_{1m} &= \sum_{\lambda_m \leq 1} \exp\{-b\lambda_m\} [E(\zeta_{1m} + \dots + \zeta_{sm} - 1 - \lambda_m)^b \\ &\quad + (-1)^b (1 + \lambda_m)^b \exp\{-\lambda_m\}] + (-1)^{b+1} \exp\{-\lambda_m\}, \end{aligned}$$

if $\lambda_m \leq 1$ then

$$E(\zeta_{1m} + \dots + \zeta_{sm} - 1 - \lambda_m)^b = \sum_{i=1}^b C_b^i (-1)^i E(\zeta_{1m} + \dots + \zeta_{sm})^{b-1} \leq C(b) \lambda_m^2 - 1 - (b-1)\lambda_m.$$

Therefore we get

$$\begin{aligned} \sum_{\lambda_m \leq 1} \Delta'''_{1m} &\leq \sum_{\lambda_m \leq 1} \exp\{-b\lambda_m\} (C(b) \lambda_m^2 - (1 + (b-1)\lambda_m)) (1 - (1 + \lambda_m) \exp\{-\lambda_m\}) \\ &\leq C(b) \sum_{m=1}^N \lambda_m^2 \exp\{-2\lambda_m\} \leq C(b) \sigma_N^2(s). \end{aligned} \quad (8)$$

Since $(1 + \lambda_m)^b \leq 1 + (b-1)\lambda_m + C(b)\lambda_m^2$, if $\lambda_m \leq 1$. Using well known inequality between moments of r.v., we obtain:

$$\begin{aligned} \sum_{\lambda_m \geq 1} \Delta'''_{1m} &\leq \sum_{\lambda_m \geq 1} \exp\{-b\lambda_m\} E|\zeta_{1m} + \dots + \zeta_{sm}|^b \\ &\leq \sum_{\lambda_m \geq 1} (E(\zeta_{1m} + \dots + \zeta_{sm} - 1 - \lambda_m)^{b+1})^{b/b+1} \exp\{-b\lambda_m\} \\ &\leq C(b) \sum_{\lambda_m \geq 1} \lambda_m^{b/2} \exp\{-b\lambda_m\} \leq C(b) \sum_{\lambda_m \geq 1} \lambda_m^2 \exp\{-2\lambda_m\} \leq C(b) \sigma_N^2(s). \end{aligned} \quad (9)$$

From (5), (6), (7), (8) and (9) follows

$$\sum_{m=1}^N \Delta_{1m} \leq C(b) \sigma_N^2(s). \quad (10)$$

Let us estimate $\sum_{m=1}^N \Delta_{2m}$. We have

$$\sum_{m=1}^N |a_k - \exp\{-\lambda_m\}|^b E|\zeta_{km} - \lambda_m|^b \leq \sum_{\lambda_{km} \leq 1} (a_k - \exp\{-\lambda_m\})^b E|\zeta_{km} - \lambda_m|^b +$$

$$\begin{aligned}
 \sum_{\lambda_{km} \geq 1} (a_k - \lambda_{km})^b E[\zeta_{km} - \lambda_{km}]^{b+1} &\leq C(b) \left[\sum_{\lambda_{km} \leq 1} (a_k - \exp\{-\lambda_m\})^2 \lambda_{km} + \right. \\
 \sum_{\lambda_{km} > 1, \alpha_k < 1} (a_k - \exp\{-\lambda_m\})^2 \lambda_{km}^{b/2} &+ \left. \sum_{\lambda_{km} > 1, \alpha_k \geq 1} |a_k - \exp\{-\lambda_m\}|^b \lambda_{km}^{b/2} \right] \leq C(b) \sigma_N^2(s) + \\
 \sum_{\lambda_{km} > 1, \alpha_k > 1} |a_k - \exp\{-\lambda_m\}|^b \lambda_{km}^{b/2}. & \quad (11)
 \end{aligned}$$

Here, we used that $\lambda_{km} \leq C_0 \alpha_k, \dots, a_k \leq 1$ and $\lambda_{km} \leq C_0$ if $\alpha_k \leq 1$, $E|\zeta_{km} - \lambda_{km}|^i \leq C(i) \lambda_{km}$.

Let $\lambda_{km} > 1, \alpha_k > 1$. Since $a_k \leq \alpha_k^{-1}$ we have

$$|a_k - \exp\{-\lambda_m\}| \lambda_m^{1/2} \leq a_k \lambda_m^{1/2} + 1 \leq \sqrt{C_0/\alpha_k} + 1 \leq \sqrt{C_0} + 1.$$

Therefore

$$\begin{aligned}
 \sum_{\lambda_{km} > 1, \alpha_k > 1} (|a_k - \exp\{-\lambda_m\}| \lambda_{km}^{1/2})^b &\leq (\sqrt{C_0} + 1)^{b-2} \sum (a_k - \exp\{-\lambda_m\})^2 \lambda_{km} \\
 &\leq (\sqrt{C_0} + 1)^{b-2} \sigma_N^2(s).
 \end{aligned}$$

From this and (11)

$$\sum_{m=1}^N \Delta_{2m} \leq C(b) \sigma_N^2(s). \quad (12)$$

Thus if b is odd, then by (4), (10), (11) it follows that

$$\sum_{m=1}^N E|g_m(c_m^{(s)})|^b \leq C(b) \sigma_N^2(s). \quad (13)$$

If b is odd then the theorem follows from (2) and (13). If b is even then the theorem follows from the well-known inequality between Ljapunov's ratio and (2), (13). Proof of theorem is complete. \square

MIRAKHMEDOV, SAIDOVA

References

- [1] Mirakhmedov, S.A., An approximations of multiple randomized divisible statistics by means of Normal distribution, *Theory Probabl. and Appl.*, **32** 761-771 (1987).
- [2] Quine, M.P., Robinson, J., A Berry-Esseen bound for an occupancy problem, *Ann. Probabl.*, **10**, 663-671, 1982.
- [3] Kolchin, V.F., Sevastjanov, B.V., Chistjakov, V.P., Random allocations, Nauka, Moskov, 1976.

Serzod A. MIRAKHMEDOV & Odila SAIDOVA
Department of Mathematics
Tashkent State University
700095, Tashkent - UZBEKISTAN

Received 13.05.1996